

Health impact assessment of exposure to fine particulate matter based on satellite and meteorological information†

Hak-Kan Lai,* Hilda Tsang, Thuan-Quoc Thach and Chit-Ming Wong

Cite this: *Environ. Sci.: Processes Impacts*, 2014, **16**, 239

Received 8th July 2013
Accepted 5th November 2013

DOI: 10.1039/c3em00357d

rsc.li/process-impacts

Air pollution in China, especially in the Pearl River Delta (PRD) region, has drastically increased in recent years. We modelled annual mean ground-level $\text{PM}_{2.5}$ concentrations based on worldwide satellite information and meteorological data from 40 cities outside the PRD. The model of $\text{PM}_{2.5}$ concentration ($R = 0.845$) was best explained by aerosol optical thickness (43.8%). We validated the spatial-temporal dimensions of the model and estimated that the annual mean $\text{PM}_{2.5}$ concentration in PRD ranged between 22 and $65 \mu\text{g m}^{-3}$. Then we used meta-analysis to obtain the pooled excess risks of mortality in China and assessed the health impacts. We found an inverse association between short-term excess risks of mortality and annual mean $\text{PM}_{2.5}$ concentrations. Based on the above models and analyses, the associated excess deaths for all-cause and cardiopulmonary diseases were 3386 and 2639 respectively. The corresponding risk-standardized excess death rates were 2006 and 1069 per million people.

Environmental impact

Particulate air pollution has surged over densely populated cities in China in recent years and has recently received international attention due to the toxicity of the fine portion of particulate matter and the associated adverse health consequences to local residents as well as to tourists and investors. The environmental impact could be assessed based on the community health burden attributable to exposure to the ambient concentration in the Pearl River Delta which represents Chinese regions with rapid economic growth where energy production mainly depends on fossil fuels with air pollution generated from power plants, motor vehicles and industrial emissions. Worldwide satellite and meteorological information could be applied for exposure assessment and model validation.

Introduction

Long-term exposure to fine particulate matter $\text{PM}_{2.5}$ (aerodynamic diameter $\leq 2.5 \mu\text{m}$) has been linearly associated with all-cause, cardiovascular and lung cancer mortality with detectable effects down to $8 \mu\text{g m}^{-3}$,¹ which is very close to the World Health Organization (WHO) Air Quality Guidelines (AQG) for $\text{PM}_{2.5}$ of $10 \mu\text{g m}^{-3}$.² Urban air pollution in Asian countries contributed to two thirds of the global burden of disease due to poor air quality³ and the highest levels were often observed in Chinese cities.⁴ China's rapid economic growth relying on vast consumption of fossil fuels⁵ for power generation,

transportation and manufacturing industries has drastically increased the emissions in recent decades.⁶ The Pearl River Delta (PRD) in south mainland China is one of the most polluted regions where $\text{PM}_{2.5}$ monitoring is still at an initial stage.⁷

Recent epidemiological studies have used correlations between satellite data of aerosol optical thickness (AOT) and $\text{PM}_{2.5}$ to estimate the spatial variations of exposure to particulate pollution^{4,8–11} and the associated adverse health impacts.¹² We aimed to estimate the $\text{PM}_{2.5}$ concentration in the PRD region based on worldwide satellite and meteorological data, to pool the excess risks of mortality from the worldwide literature, and to assess the associated excess mortality in the PRD population.

Materials and methods

$\text{PM}_{2.5}$ data

We collected $\text{PM}_{2.5}$ annual mean concentration ($\mu\text{g m}^{-3}$) for 40 cities in 25 countries (Fig. 1) corresponding to the years 2008 ($n = 31$) and 2009 ($n = 9$) from the World Health Organization (WHO) Urban Outdoor Air Pollution Database¹³ and from the Hong Kong SAR Government.¹⁴ The latter was the only city in

Department of Community Medicine, School of Public Health, The University of Hong Kong, Hong Kong SAR, China. E-mail: lai.hk@hku.hk; Fax: +852 2855 9528; Tel: +852-2819 9280

† Electronic supplementary information (ESI) available: Supplementary I: cluster analysis for selection of independent variables in multiple regressions. Supplementary II: regression models of $\text{PM}_{2.5}$ and sensitivity of the model predictions; visual inspection of the major independent variables: ln-transformed data of $\text{PM}_{2.5}$ versus AOT and visibility in 40 cities. Supplementary III: literature review – detection of heterogeneity and publication bias. Supplementary IV: summary of excess risks of different causes of mortality in all age groups due to $\text{PM}_{2.5}$. See DOI: 10.1039/c3em00357d



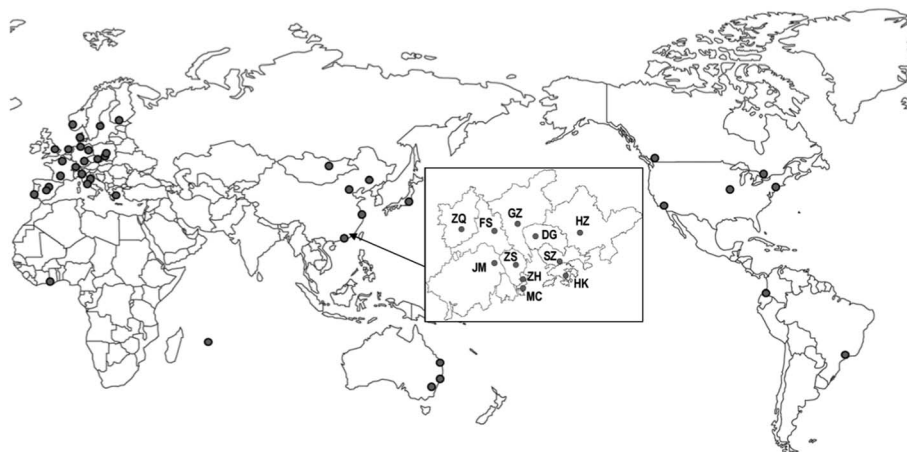


Fig. 1 Forty cities for modeling and eleven cities in the Pearl River Delta region in China for health impact assessment. Note: Large dots represent 40 cities in the world for modeling. Small dots inside the box represent 11 cities in the Pearl River Delta region in China for health impact assessment. ZQ = Zhaoqing, FS = Foshan, GZ = Guangzhou, DG = Dongguan, HZ = Huizhou, JM = Jiangmen, ZS = Zhongshan, SZ = Shenzhen, ZH = Zhuhai, MC = Macau, and HK = Hong Kong. 4 cities in China (Beijing, Shanghai, Shenyang, and Hong Kong) were part of the 40 cities in the world for modeling.

the PRD with accessible data of $PM_{2.5}$ recorded by a fixed-site monitoring network. We also obtained $PM_{2.5}$ data from Beijing, Guangzhou, Shanghai and Tokyo by contacting the authors of relevant publications.^{15,16}

Satellite and meteorological data

We retrieved the databases from National Aeronautics and Space Administration (NASA)'s Moderate-resolution Imaging Spectroradiometer (MODIS) for monthly averages of AOT,^{17,18} cloud optical thickness (COT), carbon monoxide (CO) level (ppbv), land surface temperature (LST) ($^{\circ}C$), and vegetation index (VI) in the 40 world cities as well as 11 Chinese cities in the PRD region (Fig. 1). We assessed the satellite information in 2–4 grids of fixed coordinate size (0.11×0.11) in the central area of each city (covering 272 to 547 km^2). The annual mean value for air temperature ($^{\circ}C$), dew point ($^{\circ}C$), relative humidity (%), visibility (km), wind speed ($km\ h^{-1}$), and cloud cover (okta) data of the base year were retrieved from a publicly accessible database.¹⁹ Missing data of monthly AOT (26%) were replaced by the data of the same month in the closest year (± 1 to 3 years) multiplied by the year-to-year percentage changes of the available monthly data. The remaining missing data of AOT (13%) as well as some missing data of COT and CO (1.5% and 7.3%, respectively) were replaced by a multiple imputation method.²⁰ AOT and visibility data from all cities were right-skewed and ln-transformed.

Modeling

We used cluster analysis to select independent variables with low multicollinearity and then fitted multiple regression models to estimate $PM_{2.5}$ annual mean concentrations for all combinations of these selected variables (Suppl 1†).²¹ We assessed models' temporal sensitivity by using years different from 2008 and 2009 in Hong Kong, Los Angeles, Rome, Shenyang and Sydney which were cities with non-missing data. We also assessed models' spatial sensitivity by using cities other than the

40 cities. We identified the best model as having the smallest average discrepancy in both temporal and spatial sensitivity for prediction of $PM_{2.5}$ annual mean (Suppl 2†). We ranked influential predictors based on the percentage variance explained (partial eta square η^2) by each variable in the model and then visually inspected the linear associations between $PM_{2.5}$ and the influential predictors after adjusting for covariates (Suppl 2†).²²

Systematic review

We searched the MEDLINE database using structured keyword terms, ("fine particulate"[tiab] or "fine particulates"[tiab] or "fine particle"[tiab] or "fine particles"[tiab] or "PM2.5"[tiab] or "PM(2.5)"[tiab]) and ("health") and (mortality[tiab]), and retrieved 390 abstracts on 1st June 2012. We selected 35 literature articles using the following inclusion criteria (Fig. 2a): (i) all epidemiological studies on the adverse health effects of $PM_{2.5}$ as the main research question; (ii) the health outcomes were related to deaths; (iii) the subjects were not high risk groups (e.g. patients or smokers) nor sensitive age subgroups (e.g. children or elders); (iv) exposures to ambient levels (not indoor, occupational or accidental exposures); (v) the health risk estimates were expressed in terms of unit change in pollutant mass concentration; and (vi) reported in English. We included non-duplicated publications on short-term effects only. We excluded the literature on long-term effects of mortality due to the insufficient number of studies that followed the above-mentioned criteria ($n = 5$). We included studies on the single-pollutant model with lags ranging within day 0 and 1 since these findings were mostly reported. Finally, we selected the latest study only in each of the population to avoid over-representation by any single population (Fig. 2b).

Meta-analysis

We used meta-analysis to pool risk estimates of $PM_{2.5}$ on mortality for every $10\ \mu g\ m^{-3}$ increase in pollutant



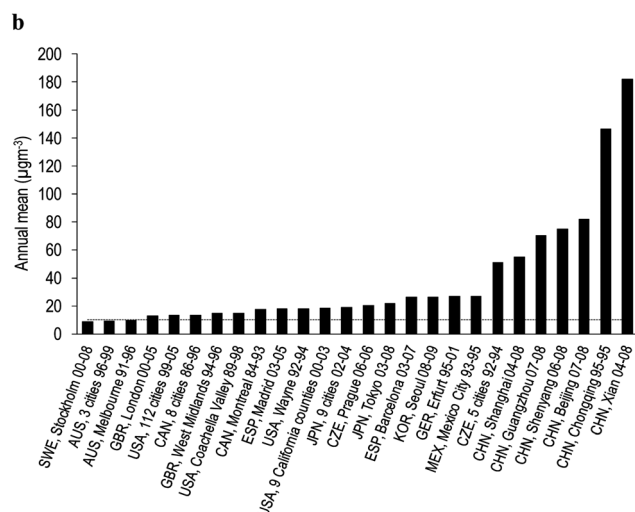
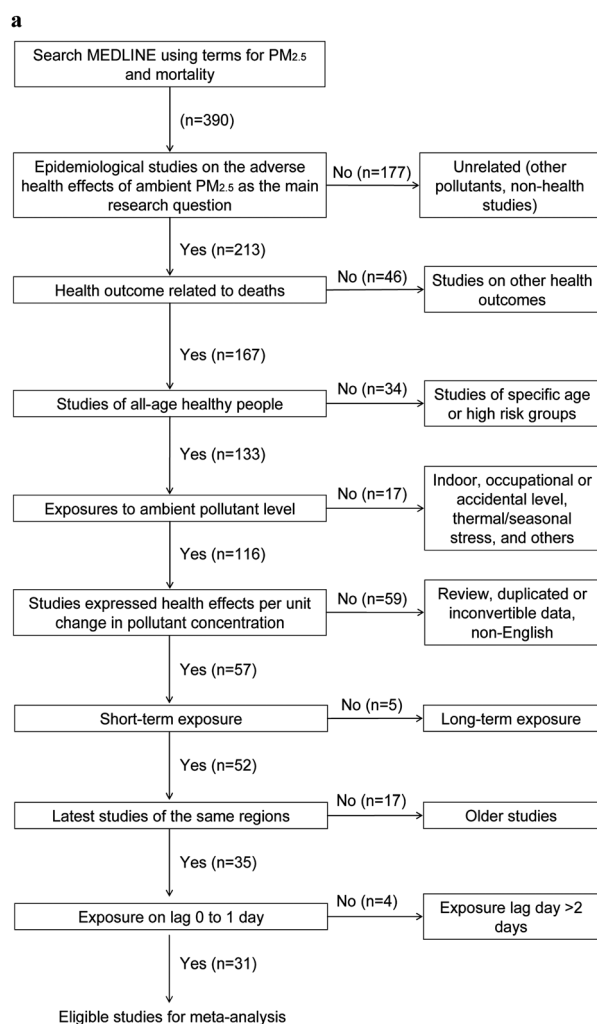


Fig. 2 (a) Literature selection. (b) Selected literature' records of PM_{2.5} concentrations. Note: Only the annual mean concentrations of the latest exposure year in each city are shown. Three-letter country codes: Australia (AUS), Canada (CAN), China (CHN), Czech Republic (CZE), Germany (GER), Japan (JPN), Korea (KOR), Mexico (MEX), Spain (ESP), Sweden (SWE), United Kingdom (GBR), and United States (USA). The dotted line represents the WHO annual AQG of 10 µg m⁻³ for PM_{2.5}.

concentration.²³ We pooled excess risks (ER) in random effects if I^2 statistics for heterogeneity was >25% or otherwise in a fixed effect. We used the influence plot of the conditional square root q^2 statistics against square root w weight to diagnose the overall heterogeneity due to each study; and funnel plots with the Egger test on asymmetry at alpha level 0.1 to assess publication bias. We assumed the overall heterogeneity as statistically significant under normal distribution when the square root of q^2 statistic lay outside -1.96 and 1.96.^{24,25}

Meta-regression

Recent reviews of the health effects of PM_{2.5} in the American,²⁶ European²⁷ and Asian countries²⁸ have consistently indicated evidence of increased deaths when exposure is elevated, but meta-analyses with meta-regression of these estimates across regions were rare. In this study, we pooled the short-term effect estimates of the excess risk of mortality from time-series studies on ambient PM_{2.5} in different global regions and conducted a meta-regression analysis of these estimates for improving our understanding of the differences in effect sizes across regions.

Health impact assessment

We obtained the annual total number of deaths (N) in 2008 from the Guangzhou Health Statistical Yearbook,²⁹ Department of Health Statistical yearbook in Macau,³⁰ and Hospital Authority clinical database in Hong Kong.³¹ The number of deaths in prefectures other than Guangzhou in PRD was not publicly available and we estimated by proportionality using the population ratios between these prefectures and Guangzhou (the capital city of Guangdong). We calculated the health impact by multiplying the pooled ER with N and the estimated annual mean concentrations of PM_{2.5}. We assumed no threshold level for PM_{2.5} based on the hypothesis of both linear³²⁻³⁵ and nonlinear concentration response relationships.^{36,37} For comparison with the global health burden assessment study of excess deaths due to PM_{2.5},³⁸ we calculated *risk-standardized excess death rates*, which was defined as the excess death rates per one million population multiplied by 10% (per 10 µg m⁻³ of PM_{2.5}) as the fixed excess risk of death.

Table 1 Selected best model of annual mean PM_{2.5} concentration (µg m⁻³) for health impact assessment^a

Model	Unit	β	95% CI	p -values	VIF	η^2 (%)
Constant		8.097	6.427, 9.767	<0.001		
ln A	0-1	0.613	0.367, 0.859	<0.001	1.5	43.8
ln V	km	-0.987	-1.587, -0.388	0.002	1.3	25.4
C	0-50	-0.048	-0.095, -0.001	0.045	1.7	11.7
W	km h ⁻¹	-0.037	-0.075, 0.002	0.063	1.1	10.1
L	°C	-0.020	-0.049, 0.009	0.165	1.8	5.7
cc	okta	-0.068	-0.233, 0.097	0.406	1.4	2.1

^a Note: A = aerosol optical thickness, V = visibility, C = cloud optical thickness, W = wind speed, L = land surface temperature, and cc = cloud cover. VIF = variance inflation factor, in which VIF less than 2.5 is indicative of non-collinearity. η^2 = partial eta square, which indicates the percentage variance explained by the variable in the model. Model selection (see M8 in Suppl 2, ESI).



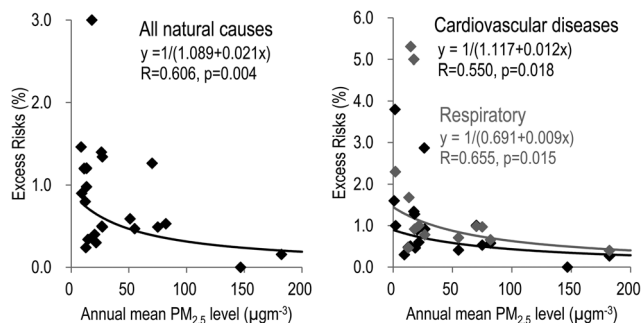


Fig. 3 Meta-regression analysis of annual mean $\text{PM}_{2.5}$ concentration and the non-negative excess risk (%) of mortality for all natural causes ($n = 22$) and for cardiovascular ($n = 19$) and respiratory causes ($n = 13$) in different studies.

Results

Table 1 shows the selected best model ($R = 0.85$) that $\text{PM}_{2.5}$ concentration was positively associated with \ln -transformed aerosol optical thickness ($\beta = 0.613$ [95%CI: 0.367, 0.859]) but

negatively associated with \ln -transformed visibility ($\beta = -0.987$ [−1.587, −0.388]). These associations were consistent in other models. The largest variance of $\text{PM}_{2.5}$ was explained by \ln -transformed AOT (44%, $p < 0.001$) and multi-collinearity (variance inflation factors: 1.1 to 1.8) which were low for the rest of the covariates.

The pooled ER of mortality due to all-cause (ICD-10 A00–R99), cardiovascular (ICD-10 I00–I99) and respiratory (ICD-10 J00–J99) diseases for all the reviewed studies were 0.69% [0.45, 0.94], 0.60% [0.39, 0.80] and 0.83% [0.50, 1.17] respectively (Suppl 3 and 4†).^{63–93} The highest pooled ER by region in America (1.01% [0.79, 1.22]) followed by Europe 0.88% [0.34, 1.42] and Asia-Pacific 0.50% [0.19, 0.81] while the pooled ER for all of the six Chinese cities was 0.51% [0.14, 0.89] (Suppl 3†). Some publication bias was found (Egger test: $p = 0.080$) but it was not evident (Egger test: $p = 0.460$) after excluding influential studies defined by significant overall heterogeneity (Suppl 5†). Results of meta-regression analysis indicated that the ER of mortality in the literature were inversely related to the annual $\text{PM}_{2.5}$ concentrations in a slightly decaying pattern (Fig. 3).

Table 2 Health impact assessment – excess deaths due to $\text{PM}_{2.5}$ in the PRD regions^a

	HK	MC	DG	FS	GZ	HZ	JM	SZ	ZH	ZQ	ZS	PRD
Annual mean ($\mu\text{g m}^{-3}$)												Mean
Monitored $\text{PM}_{2.5}$	33											45.2
Modeled $\text{PM}_{2.5}$		28	58	65	61	43	32	47	22	63	47	
Estimated $\text{PM}_{2.5}$ (70% of monitored PM_{10} , for validation purpose)	33	39	56	81	50	55	48	38	39	62	49	50.0
												Total
Population (million)	7.0	0.5	1.7	3.6	7.8	3.2	3.9	2.3	1.0	4.1	1.5	36.7
Death rate per 1000	6.00	3.20	4.54	5.49	5.55	4.90	7.49	0.98	2.79	5.25	6.01	
												Total
No. of deaths (1000)												
All-cause	39.8	1.8	7.9	20.0	43.5	15.6	28.9	2.3	2.8	21.5	8.8	192.9
Cardiovascular	11.4	0.5	2.8	7.1	15.5	5.6	10.3	0.8	1.0	7.7	3.1	65.9
Respiratory	8.2	0.2	1.4	3.6	7.9	2.8	5.3	0.4	0.5	3.9	1.6	36.0
												Mean
ER derived from meta-regressions												
All-cause	0.56	0.60	0.43	0.41	0.42	0.50	0.57	0.48	0.64	0.42	0.48	0.49
Cardiovascular	0.66	0.69	0.55	0.53	0.54	0.61	0.67	0.60	0.72	0.53	0.59	0.60
Respiratory	1.02	1.06	0.83	0.79	0.81	0.93	1.02	0.90	1.12	0.80	0.90	0.91
												Total
ED												
All-cause	505	19	165	447	937	259	362	40	22	472	157	3386
Cardiovascular	170	6	75	206	429	113	152	18	9	217	70	1463
Respiratory	187	4	57	157	326	87	119	14	7	165	53	1176
												Total
ED rate (per million)												
All-cause	72	34	94	123	120	81	94	17	22	115	108	92
Cardiovascular	24	11	43	57	55	35	39	8	9	53	47	40
Respiratory	27	8	33	43	42	27	31	6	7	40	37	32
												Total
Risk-standardized ED rate (per million)												
All-cause	1282	562	2175	3002	2835	1616	1658	359	339	2772	2238	2006
Cardiovascular	366	155	777	1072	1012	577	592	128	121	990	799	698
Respiratory	263	77	396	547	517	294	302	65	62	505	408	371

^a Note: $\text{PM}_{2.5}$ in Hong Kong was the real data from the Hong Kong Environmental Protection Department. HK = Hong Kong, MC = Macau, DG = Dongguan, FS = Foshan, GZ = Guangzhou, HZ = Huizhou, JM = Jiangmen, SZ = Shenzhen, ZH = Zhuhai, ZQ = Zhaoqing, ZS = Zhongshan, and PRD = Pearl River Delta. ER = excess risks per $10 \mu\text{g m}^{-3}$ derived from meta-regression models (see Fig. 3). ED = excess deaths. The risk-standardized ED rate was calculated using 10% as the excess risks per $10 \mu\text{g m}^{-3}$ of $\text{PM}_{2.5}$. The ratio of $\text{PM}_{2.5}$ to PM_{10} annual mean has been very stable at 0.7 over the years in Hong Kong, which is the only city in PRD with a fixed site monitoring of $\text{PM}_{2.5}$.



In PRD regions in 2008–2009, the model predicted annual mean $\text{PM}_{2.5}$ concentrations were $45 \mu\text{g m}^{-3}$, which was 4.5 times of the WHOAQG of $10 \mu\text{g m}^{-3}$. The lowest predicted annual mean was in Zhuhai ($22 \mu\text{g m}^{-3}$) and the highest one was in Foshan ($65 \mu\text{g m}^{-3}$). In the region, the mean estimated $\text{PM}_{2.5}$ annual concentration based on the ratio of $\text{PM}_{2.5}/\text{PM}_{10}$ was $50 \mu\text{g m}^{-3}$, with the lowest annual mean in Hong Kong ($33 \mu\text{g m}^{-3}$) and the highest one in Foshan ($81 \mu\text{g m}^{-3}$) (Table 2). The predicted values of the $\text{PM}_{2.5}$ in PRD regions were relatively higher than the rest of the other cities in the world (Fig. 4).

Based on model predicted $\text{PM}_{2.5}$ levels, the annual total number of deaths due to all natural causes ranged from 1756 in Macau to 43 521 in Guangzhou, with a total of 192 911 deaths per year in the whole region (Table 2). The annual cardiovascular deaths were the lowest in Macau (485) and the highest in Guangzhou (15 543); but the highest respiratory deaths were observed in Hong Kong (8160).

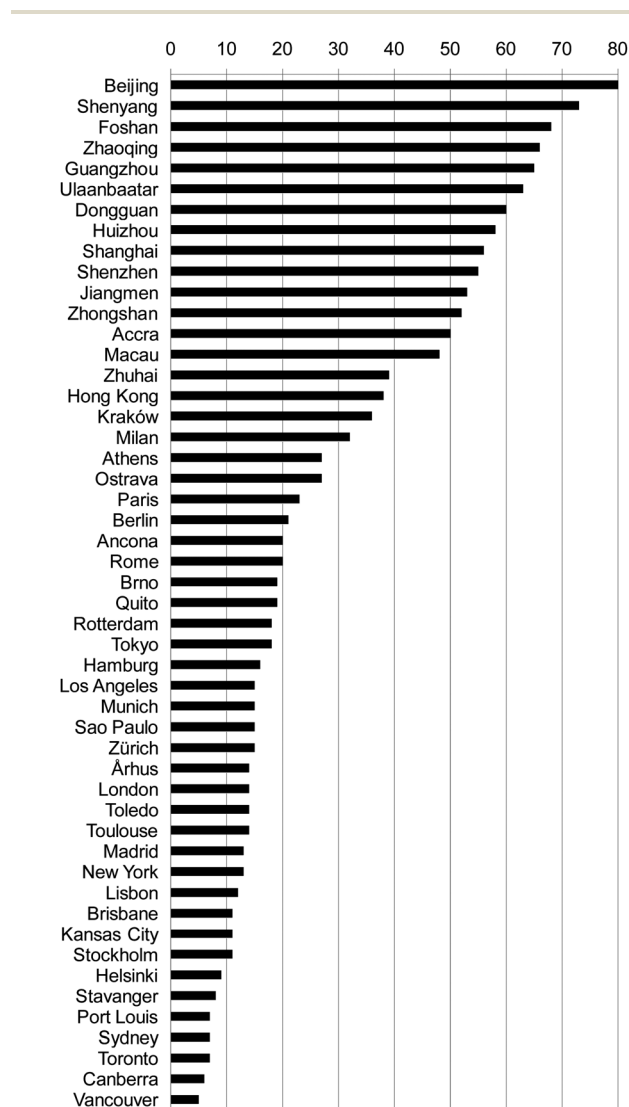


Fig. 4 Modeled and monitored annual mean $\text{PM}_{2.5}$ concentrations ($\mu\text{g m}^{-3}$) in all cities from 2008–2009.

The annual excess deaths due to model-predicted $\text{PM}_{2.5}$ for all causes were the highest in Guangzhou (937) and the lowest in Macau (19). The same trends followed in excess deaths due to cardiovascular and respiratory diseases. The total excess deaths due to all-cause, cardiovascular and respiratory diseases were 3386, 1463, and 1176 respectively in the whole PRD region in 2008–2009. The corresponding risk-standardized excess death rates per one million people were 2006, 698 and 371 respectively (Table 2).

Discussion

Many studies have examined the relationship between ground-level monitored $\text{PM}_{2.5}$ concentration and satellite AOT on a local or countrywide scale^{10,11,39–50} as well as on a global scale.^{4,51} We based our study on the established approach of predicting ground-level $\text{PM}_{2.5}$ annual mean concentration in the PRD region where public accessibility to the monitoring data was not yet fully available.

Our model utilized both the vertical column aerosol, as measured by AOT, and the horizontal ground-level aerosol, as measured by visibility, together with cloud optical thickness and other satellite and meteorological information as co-predictors of the ground-level annual mean of $\text{PM}_{2.5}$ concentration. Visibility is a well-known indicator of $\text{PM}_{2.5}$.^{52,53} It is measured horizontally at a height that is relevant to the ground level monitoring of $\text{PM}_{2.5}$ and provides information to explain the variance of $\text{PM}_{2.5}$ in addition to that measured vertically by AOT. Besides, visibility data are readily obtainable in many cities for worldwide comparisons as they are standard measurements in airports as well as in weather observatories. Cloud optical thickness is a satellite measure of light extinction due to the amount and thickness of a cloud, which is related to humidity and pressure in the atmosphere. This variable together with cloud cover in our model shared the variance of relative humidity as a common covariate.¹⁰ Although the satellite detected carbon monoxide was identified as a covariate of AOT, it was not included in our selected model. Further studies on applying the carbon monoxide variable should be warranted since it is publicly available and is related to fossil fuel burning and traffic emissions, and we found the best predictive performance in the temporal sensitivity assessment.

It is interesting that the annual mean concentrations in the Chinese cities were relatively higher than all other studies, but the magnitude of pooled ER of all natural mortalities for these six Chinese cities was only a half of that in the American region. The APHENA investigators have also tested whether the concentration–response curve was actually nonlinear and whether statistical methods were appropriate though conclusive evidence was not explicit.⁵⁴ A distributed lag nonlinear model was also proposed to detect the nonlinear concentration–response curve and delayed effects.⁵⁵ In meta-regression analysis we found some evidence of an inverse linear relationship between ER of mortality and the annual $\text{PM}_{2.5}$ concentration. This pattern appears to be leveling off in the downward direction rather than purely linear.³⁶ A similar pattern of inverse association has been observed in a cross-sectional study of lung



function and exposure to indoor PM_{2.5} concentration.³⁷ This inverse relationship could be related to a saturation mechanism occurring at lower exposure levels⁵⁶ where both irreversible and reversible processes may simultaneously exist. This has been hypothesized that, as indicated by structural changes in airways,^{57,58} particulates that penetrate to the deepest part of the lung and cause alveolar epithelial injury are associated with both acute reversible inflammatory responses and cumulative irreversible pulmonary damage.³⁷ An explanation by the saturation hypothesis for cardiovascular diseases is also warranted.³⁶ Our findings suggest that there is a need for further study to assess the effects of longer-term exposure on the short-term mortality risks. In addition, difference in climate, population structures, gene susceptibility, health status, occupational exposures, long-term exposure history and composition in PM_{2.5} may play a role in explaining the potential differences in short-term effects by region. There were some reviews on the epidemiological studies on the impact of fine particulate air pollution on mortality,^{26,59,60} while reviews on the cause for specific mortality are rare. The present review provided a meta-regression between the short-term effects and the annual mean concentrations and is the first to report an inverse relationship with a slightly decaying pattern.

For health impact assessment in the PRD region, we reviewed six Chinese studies and pooled the excess risks of mortalities based on physician-diagnosis, which were once-in-a-lifetime clinical records and so could provide more reliable health effect information than physiological and bio-microscopic measurements. Our health impact assessment results were well supported by the latest global estimation of health burden due to PM_{2.5} concentration³⁸ since our estimation of the risk-standardized excess death rate of 1069 per million people for cardiorespiratory causes was very close to that in Asia of 1104 per million per 10% excess risk (*i.e.* 2584 thousand \times 1 million/1.8 billion \times 10 excess risk%/13 reported excess risk%) in that report. Non-risk standardized estimates of disease burden varies between places and the type of study design. For instance, the short-term excess death rates calculated in the current study were smaller by an order of magnitude than those estimated in long-term studies.^{34,59} However when we standardized the excess risk to 10%, comparison with other studies was more meaningful.

There were several limitations in this study. First, the number of cities involved in modeling was based on availability of satellite AOT data and ground-level fixed site monitored records of PM_{2.5} concentrations as compiled by WHO.⁵ Also some cities might have been overlooked, but the approximate normal distributions of our collected data have indicated that the cities selected in our sample are representative of respective regions so that selection bias might have been minimum and therefore not affecting overall representativeness of the sample. Second, our assumption of PM_{2.5} concentration as 70% of PM₁₀ concentration in the PRD region might have ignored the heterogeneity of PM_{2.5}/PM₁₀ ratio across the region. However a similar ratio derived from a recent study in Guangzhou has indicated that it is in good agreement with the ratio we used.⁶¹ Third, the generalization of our best model for prediction in

cities outside China has not been assessed despite that our spatial validation was based on ten cities in the PRD region only. Nevertheless, our temporal validation that was based on cities in different continents has demonstrated to have good performance for prediction among the forty cities that were included in the modeling. Fourth, the variation in risks due to differences in anthropogenic *versus* natural dust components of PM was not addressed.⁶² Our results should be interpreted with caution especially for international comparison.

Conclusions

The exposure to particulate pollutants in the Pearl River Delta regions in China is much higher than many other cities in the world. This study shows that the variation in PM_{2.5} annual concentrations may be inversely related to magnitude of short-term effects on mortality across geographic regions. The rapid growth of the economy in the region undoubtedly has contributed to the impact of air pollution on the population which is estimated to be responsible for 1069 excess deaths per million people. Nowadays, satellite and meteorological information is readily accessible so that public health impacts due to particulate air pollution in locations lacking monitoring can also be assessed.

Acknowledgements

This work was supported by the Wellcome Trust [#094330/Z10/Z].

Notes and references

- 1 J. Lepeule, F. Laden, D. Dockery and J. Schwartz, *Environ. Health Perspect.*, 2012, **120**, 165–170.
- 2 World Health Organization, *Air Quality Guidelines Global Update 2005: particulate matter, ozone, nitrogen dioxide and sulfur dioxide*. WHO Regional Office for Europe, Copenhagen, 2006.
- 3 M. Krzyzanowski and A. Cohen, *Air Qual., Atmos. Health*, 2008, **1**, 7–13.
- 4 A. van Donkelaar, R. V. Martin, M. Brauer, R. Kahn, R. Levy, C. Verduzco and P. J. Villeneuve, *Environ. Health Perspect.*, 2010, **118**, 847–855.
- 5 J. N. Wang, Y. Lei, J. T. Yang and G. Yan, *Environ. Sci. Technol.*, 2012, **46**, 4263–4264.
- 6 B. H. Chen, H. D. Kan, R. J. Chen, S. H. Jiang and C. J. Hong, *J. Air Waste Manage. Assoc.*, 2011, **61**, 1292–1299.
- 7 Y. Yuan, S. S. Liu, R. Castro and X. B. Pan, *Environ. Sci. Technol.*, 2012, **46**, 3627–3628.
- 8 J. Wang and S. A. Christopher, *Geophys. Res. Lett.*, 2003, **30**.
- 9 J. A. Engel-Cox, R. M. Hoff and A. D. J. Haymet, *J. Air Waste Manage. Assoc.*, 2004, **54**, 1360–1371.
- 10 Y. Liu, J. A. Sarnat, A. Kilaru, D. J. Jacob and P. Koutrakis, *Environ. Sci. Technol.*, 2005, **39**, 3269–3278.
- 11 Y. Liu, C. J. Paciorek and P. Koutrakis, *Environ Health Persp.*, 2009, **117**, 886–892.
- 12 D. L. Crouse, P. A. Peters, A. van Donkelaar, M. S. Goldberg, P. J. Villeneuve, O. Brion, S. Khan, D. O. Atari, M. Jerrett,



- C. A. Pope, M. Brauer, J. R. Brook, R. V. Martin, D. Stieb and R. T. Burnett, *Environ. Health Perspect.*, 2012, **120**, 708–714.
- 13 World Health Organization, Urban Outdoor Air Pollution Database, WHO, Geneva, 2011. http://www.who.int/phe/health_topics/outdoorair/databases/en/index.html Last accessed 13 June 2012.
 - 14 Hong Kong Environmental Protection Department. Past air monitoring station data, 2008–2009. <http://epic.epd.gov.hk/ca/uid/airdata/p/1> Last accessed 21 May 2012.
 - 15 R. Chen, Y. Li, Y. Ma, G. Pan, G. Zeng, X. Xu, B. Chen and H. Kan, *Sci. Total Environ.*, 2011, **409**, 4934–4938.
 - 16 T. Yorifuji, I. Kawachi, M. Kaneda, S. Takao, S. Kashima and H. Doi, *Sci. Total Environ.*, 2011, **409**, 3620–3627.
 - 17 H. K. Lai, S. Y. Ho, C. M. Wong, K. K. Mak, W. S. Lo and T. H. Lam, *Int. J. Environ. Health Res.*, 2010, **20**, 219–230.
 - 18 Y. Liu, M. Franklin, R. Kahn and P. Koutrakis, *Remote Sensing of Environment*, 2007, **107**, 33–44.
 - 19 Weather underground. Weather History for cities in 2008 and 2009. <http://www.wunderground.com>, 2013 Weather Underground, Inc, Last accessed 1 May 2012.
 - 20 J. L. Schafer and J. W. Graham, *Psychological Methods*, 2002, **7**, 147–177.
 - 21 H. K. Lai, L. Bayer-Oglesby, R. Colville, T. Gotschi, M. J. Jantunen, N. Künzli, E. Kulinskaya, C. Schweizer and M. J. Nieuwenhuijsen, *Atmos. Environ.*, 2006, **40**, 1299–1313.
 - 22 W. A. Larsen and S. J. McCleary, *Technometrics*, 1972, **14**, 781–790.
 - 23 H. K. Lai, H. Tsang and C. M. Wong, *BMC Public Health*, 2013, **13**, 360.
 - 24 M. Woodward. *Meta-analysis (Ch.12). Epidemiology: study design and data analysis*, 2nd edn, Fla: Chapman & Hall/CRC, Boca Raton, 2005.
 - 25 M. Egger, S. G. Davey and M. C. Schneider, *BMJ*, 1997, **315**, 629–634.
 - 26 J. I. Levy, D. Diez, Y. Dou, C. D. Barr and F. Dominici, *Am. J. Epidemiol.*, 2012, **175**, 1091–1099.
 - 27 H. R. Anderson, R. W. Atkinson, J. L. Peacock, L. Marston and K. Konstantinou. *Meta-analysis of time-series studies and panel studies of particulate matter (PM) and ozone (O3): report of a WHO task group*, World Health Organization, Copenhagen, 2004.
 - 28 H. J. Jahn, A. Schneider, S. Breitner, R. Eissner, M. Wendisch, A. Kramer, *et al.*, *Int. J. Hyg. Environ. Health*, 2011, **214**, 281–295.
 - 29 Bureau of Health of Guangzhou Municipality, 2010.
 - 30 Macao Department of Health, Boletim Estatístico, 2008. http://www.ssm.gov.mo/design/statistic/e_statistic_fs.htm, Last accessed 2 February 2012.
 - 31 Hong Kong Hospital Authority. Hospital Authority Statistical Report, 2009, http://www.ha.org.hk/upload/publication_15/281.pdf, Last accessed 2 February 2012.
 - 32 M. J. Daniels, F. Dominici, J. M. Samet and S. L. Zeger, *American journal of epidemiology*, 2000, **152**, 397–406.
 - 33 E. Samoli, A. Analitis, G. Touloumi, J. Schwartz, H. R. Anderson, J. Sunyer, L. Bisanti, D. Zmirou, J. M. Vonk, J. Pekkanen, P. Goodman, A. Paldy, C. Schindler and K. Katsouyanni, *Environ. Health Perspect.*, 2005, **113**, 88–95.
 - 34 C. A. Pope 3rd and D. W. Dockery, *J. Air Waste Manage. Assoc.*, 2006, **56**, 709–742.
 - 35 C. M. Wong, N. Vichit-Vadakan, H. Kan and Z. Qian, *Environ. Health Perspect.*, 2008, **116**, 1195–1202.
 - 36 K. R. Smith and J. L. Peel, *Environ Health Persp*, 2010, **118**, 1643–1645.
 - 37 H. K. Lai, A. J. Hedley, J. Repace, C. So, Q. Y. Lu, S. M. McGhee, R. Fielding and C. M. Wong, *Thorax*, 2011, **66**, 615–623.
 - 38 S. C. Anenberg, L. W. Horowitz, D. Q. Tong and J. J. West, *Environ Health Persp*, 2010, **118**, 1189–1195.
 - 39 A. van Donkelaar, R. V. Martin and R. J. Park, *J. Geophys. Res.: Atmos.*, 2006, 111.
 - 40 N. Kumar, A. Chu and A. Foster, *Atmos. Environ.*, 2007, **41**, 4492–4503.
 - 41 Z. F. Wang, L. F. Chen, J. H. Tao, Y. Zhang and L. Su, *Remote Sensing of Environment*, 2010, **114**, 50–63.
 - 42 P. Glantz, A. Kokhanovsky, W. von Hoyningen-Huene and C. Johansson, *Atmos. Environ.*, 2009, **43**, 5838–5846.
 - 43 C. J. Paciorek and Y. Liu, *Environ. Health Perspect.*, 2009, **117**, 904–909.
 - 44 Y. Liu, P. Koutrakis and R. Kahn, *J. Air Waste Manage. Assoc.*, 2007, **57**, 1351–1359.
 - 45 Y. Liu, P. Koutrakis, R. Kahn, S. Turquety and R. M. Yantosca, *J. Air Waste Manage. Assoc.*, 2007, **57**, 1360–1369.
 - 46 C. J. Paciorek, Y. Liu, H. Moreno-Macias and S. Kondragunta, *Environ. Sci. Technol.*, 2008, **42**, 5800–5806.
 - 47 J. Tian and D. M. Chen, *Remote Sensing of Environment*, 2010, **114**, 221–229.
 - 48 M. Green, S. Kondragunta, P. Ciren and C. Xu, *J. Air Waste Manage. Assoc.*, 2009, **59**, 1082–1091.
 - 49 C. Li, N. C. Hsu and S. C. Tsay, *Atmos. Environ.*, 2011, **45**, 3663–3675.
 - 50 W. Di Nicolantonio and A. Cacciari, *Ital. J. Remote Sens.*, 2011, **43**, 97–109.
 - 51 P. Gupta, S. A. Christopher, J. Wang, R. Gehrig, Y. Lee and N. Kumar, *Atmos. Environ.*, 2006, **40**, 5880–5892.
 - 52 J. L. Wang, Y. H. Zhang, M. Shao, X. L. Liu, L. M. Zeng, C. L. Cheng and X. F. Xu, *J. Environ. Sci.*, 2006, **18**, 475–481.
 - 53 D. E. Abbey, B. E. Ostro, G. Fraser, T. Vancuren and R. J. Burchette, *J. Exposure Sci. Environ. Epidemiol.*, 1995, **5**, 161–180.
 - 54 K. Katsouyanni, J. M. Samet, H. R. Anderson, R. Atkinson, A. Le Tertre, S. Medina, *et al.*, *Res Rep Health Eff Inst*, 2009, 5–90.
 - 55 A. Gasparrini, B. Armstrong and M. G. Kenward, *Stat. Med.*, 2010, **29**, 2224–2234.
 - 56 G. Liistro, C. Vanwelde, W. Vincken, J. Vandevorode, G. Verleden and J. Buffels, *Chest*, 2006, **130**, 657–665.
 - 57 D. W. Dockery, F. E. Speizer, B. G. Ferris, J. H. Ware, T. A. Louis and A. Spiro, *Am. Rev. Respir. Dis.*, 1988, **137**, 286–292.
 - 58 A. D. Flouris, G. S. Metsios, A. E. Carrillo, A. Z. Jamurtas, K. Gourgoulisanis, T. Kiropoulos, M. N. Tzatzarakis,



- A. M. Tsatsakis and Y. Koutedakis, *Am. J. Respir. Crit. Care Med.*, 2009, **179**, 1029–1033.
- 59 N. Künzli, S. Medina, R. Kaiser, P. Quenel, F. Horak Jr and M. Studnicka, *Am. J. Epidemiol.*, 2001, **153**, 1050–1055.
- 60 M. L. Bell, J. M. Samet and F. Dominici, *Annu. Rev. Public Health*, 2004, **25**, 247–328.
- 61 C. Yang, X. Peng, W. Huang, R. Chen, Z. Xu, B. Chen and H. Kan, *Int Arch Occup Environ Health*, 2012, **85**, 579–585.
- 62 J. Evans, A. van Donkelaar, R. V. Martin, R. Burnett, D. G. Rainham, N. J. Birkett and D. Krewski, *Environ. Res.*, 2013, **120**, 33–42.
- 63 R. T. Burnett, J. Brook, T. Dann, C. Delocla, O. Philips and S. Cakmak, *Inhal Toxicol*, 2000, **12**, 15–39.
- 64 V. H. Borja-Aburto, M. Castillejos, D. R. Gold, S. Bierzwinski and D. Loomis, *Environ. Health Perspect.*, 1998, **106**, 849–855.
- 65 A. Zanobetti and J. Schwartz, *Environ. Health Perspect.*, 2009, **117**, 898–903.
- 66 B. D. Ostro, R. Broadwin and M. J. Lipsett, *J. Exposure Anal. Environ. Epidemiol.*, 2000, **10**, 412–419.
- 67 J. Schwartz, D. W. Dockery and L. M. Neas, *J. Air Waste Manage. Assoc.*, 1996, **46**, 927–939.
- 68 A. Peters, J. Skorkovsky, F. Kotesovec, J. Brynda, C. Spix, H. E. Wichmann, *et al.*, *Environ. Health Perspect.*, 2000, **108**, 283–287.
- 69 M. Branis, J. Vyskovska, M. Maly and J. Hovorka, *Inhalation Toxicol.*, 2010, **22**(2), 21–28.
- 70 A. Peters, S. Breitner, J. Cyrys, M. Stolzel, M. Pitz, G. Wolke, *et al.*, *Res Rep Health Eff Inst*, 2009, 5–77, 79–90, discussion.
- 71 B. Ostro, A. Tobias, X. Querol, A. Alastuey, F. Amato, J. Pey, *et al.*, *Environ. Health Perspect.*, 2011, **119**, 1781–1787.
- 72 A. Tobias, L. Perez, J. Diaz, C. Linares, J. Pey, A. Alastruey, *et al.*, *Sci. Total Environ.*, 2011, **412–413**, 386–389.
- 73 K. Meister, C. Johansson and B. Forsberg, *Environ. Health Perspect.*, 2012, **120**, 431–436.
- 74 R. W. Atkinson, G. W. Fuller, H. R. Anderson, R. M. Harrison and B. Armstrong, *Epidemiology*, 2010, **21**, 501–511.
- 75 H. R. Anderson, S. A. Bremner, R. W. Atkinson, R. M. Harrison and S. Walters, *ccup. Environ. Med.*, 2001, **58**, 504–510.
- 76 R. Simpson, G. Williams, A. Petroeschovsky, T. Best, G. Morgan, L. Denison, *et al.*, *Aust N Z J Public Health*, 2005, **29**, 205–212.
- 77 R. Chen, Y. Li, Y. Ma, G. Pan, G. Zeng, X. Xu, B. Chen and H. Kan, *Sci. Total Environ.*, 2011, **409**, 4934–4938.
- 78 S. A. Venners, B. Wang, Z. Xu, Y. Schlatter, L. Wang and X. Xu, *Environ. Health Perspect.*, 2003, **111**, 562–567.
- 79 B. H. Chen, H. D. Kan, R. J. Chen, S. H. Jiang and C. J. Hong, *J. Air Waste Manage. Assoc.*, 2011, **61**, 1292–1299.
- 80 Y. Ma, R. Chen, G. Pan, X. Xu, W. Song, B. Chen, *et al.*, *Sci. Total Environ.*, 2011, **409**, 2473–2477.
- 81 J. Cao, H. Xu, Q. Xu, B. Chen and H. Kan, *Environ. Health Perspect.*, 2012, **120**, 373–378.
- 82 J. Y. Son, J. T. Lee, K. H. Kim, K. Jung and M. L. Bell, *Environ. Health Perspect.*, 2012, **120**, 872–878.
- 83 R. Simpson, L. Denison, A. Petroeschovsky, L. Thalib and G. Williams, *J. Exposure Anal. Environ. Epidemiol.*, 2000, **10**, 488–496.
- 84 G. E. Morgan, S. Corbett, J. Wlodarczyk and P. Lewis, *Am. J. Public Health*, 1998, **88**, 759–764.
- 85 M. S. Goldberg, R. T. Burnett, J. C. Bailar, J. Brook, Y. Bonvalot, R. Tamblin, *et al.*, *Environ. Res.*, 2001, **86**, 26–36.
- 86 B. Ostro, W. Y. Feng, R. Broadwin, S. Green and M. Lipsett, *Environ. Health Perspect.*, 2007, **115**, 13–19.
- 87 W. E. Wilson, T. F. Mar and J. Q. Koenig, *J. Exposure Anal. Environ. Epidemiol.*, 2007, **17**(2), S11–S19.
- 88 J. C. Slaughter, E. Kim, L. Sheppard, J. H. Sullivan, T. V. Larson and C. Claiborn, *J. Exposure Anal. Environ. Epidemiol.*, 2005, **15**, 153–159.
- 89 M. Lippmann, K. Ito, A. Nadas and R. T. Burnett, *Res Rep Health Eff Inst*, 2000, 5–72, 73–82, discussion.
- 90 M. S. Goldberg, J. C. Bailar, R. T. Burnett, J. R. Brook, R. Tamblin, Y. Bonvalot, *et al.*, *Res Rep Health Eff Inst*, 2000, 7–113, 115–120, discussion.
- 91 B. Ostro, R. Broadwin, S. Green, W. Y. Feng and M. Lipsett, *Environ. Health Perspect.*, 2006, **114**, 29–33.
- 92 K. Ueda, H. Nitta and M. Ono, *Circulation*, 2009, **73**, 1248–1254.
- 93 M. Stolzel, S. Breitner, J. Cyrys, M. Pitz, G. Wolke, W. Kreyling, *et al.*, *J. Exposure Anal. Environ. Epidemiol.*, 2007, **17**, 458–467.

